## 18.600 Recitation 3 Recitation Instructor: Vishesh Jain math.mit.edu/~visheshj Thursday, Sep. 27th, 2018

**Solution Problem 1.** Let's start by computing the pmf of X. For  $0 \le i \le 4$ , we have

$$\mathbb{P}\left[X=i\right] = \frac{\binom{4}{i}}{16}.$$

Therefore,

$$\mathbb{E}[Y] = \sum_{i=0}^{4} \left( i(i-2)^2(i-4) \right) \frac{\binom{4}{i}}{16}$$
  
=  $\left( 1(1-2)^2(1-4) \right) \frac{\binom{4}{1}}{16} + \left( 3(3-2)^2(3-4) \right) \frac{\binom{4}{3}}{16}$   
=  $-\frac{12}{16} + -\frac{12}{16}$   
=  $-\frac{3}{2}$ .

Next,

$$\mathbb{E}\left[Y^{4}\right] = \sum_{i=0}^{4} \left(i(i-2)^{2}(i-4)\right)^{4} \frac{\binom{4}{i}}{16}$$
$$= 81 \times \frac{4}{16} + 81 \times \frac{4}{16}$$
$$= \frac{81}{4}.$$

Note that this is *not* equal to  $\mathbb{E}[Y]^4$ .

Solution Problem 2. Method 1: Direct computation using the probability mass function.

Let  $X_i$  be the random variable which is equal to 1 if the fan makes the  $i^{th}$  throw and is 0 if the fan misses the  $i^{th}$  throw. Let W be the amount of money (in dollars) won by the fan. Then,

$$W = 500(X_1 + X_2 + X_3).$$

Since the throws are independent, we have

$$\begin{split} \mathbb{P}[W = 0] &= .8 \times .8 \times .5 = .32 \\ \mathbb{P}[W = 500] &= .2 \times .8 \times .5 + .8 \times .2 \times .5 + .8 \times .8 \times .5 = .48 \\ \mathbb{P}[W = 1000] &= .2 \times .2 \times .5 + .8 \times .2 \times .5 + .2 \times .8 \times .5 = .18 \\ \mathbb{P}[W = 1500] &= .2 \times .2 \times .5 = .02 \end{split}$$

Therefore, we get

$$\mathbb{E}[W] = .48 \times 500 + .18 \times 1000 + .02 \times 1500 = 450.$$

Method 2: Use linearity of expectation.

We have

$$\mathbb{E}[W] = 500\mathbb{E}[X_1 + X_2 + X_3]$$

$$= 500 (\mathbb{E}[X_1] + \mathbb{E}[X_2] + \mathbb{E}[X_3])$$
  
= 500 (.2 + .2 + .5)  
= 500 (.9)  
= 450.

**Solution Problem 3.** Let  $D_1$  and  $D_2$  denote the (independent) random variables recording the outcomes of the two fair dice. Then,  $X = D_1 + D_2$ .

(i) To compute the pmf of X, note that there are 36 outcomes, all of which are equally probable. Moreover, for  $1 \le i \le 6$ , the number of outcomes for which  $X(=D_1 + D_2) = i$  is exactly the number of solutions to the equation  $x_1 + x_2 = i$ , where  $6 \ge x_1, x_2 \ge 1$  are integers. If  $2 \le i \le 6$ , there are (i - 1) such solutions. On the other hand, if  $7 \le i \le 12$ , there are (13 - i) such solutions. Hence, we see that

$$\mathbb{P}[X=i] = \begin{cases} i-1 & 2 \le i \le 6\\ 13-i & 7 \le i \le 12 \end{cases}.$$

Therefore,

$$\mathbb{E}[X] = \sum_{i=2}^{6} \frac{i(i-1)}{36} + \sum_{i=7}^{12} \frac{i(13-i)}{36}$$
$$= \sum_{i=2}^{6} \frac{i^2 - i}{36} + \sum_{i=7}^{12} \frac{13i - i^2}{36}$$
$$= \frac{70 + 182}{36}$$
$$= 7.$$

(ii) The problem becomes *much* easier using linearity of expectation. Note that  $\mathbb{E}[D_1] = \mathbb{E}[D_2] = \frac{1}{6} \sum_{i=1}^{6} i = \frac{7}{2}$ . Therefore,

$$\mathbb{E}[X] = \mathbb{E}[D_1 + D_2]$$
  
=  $\mathbb{E}[D_1] + \mathbb{E}[D_2]$   
=  $2 \times (7/2)$   
= 7.

(iii) Since we already know the pmf of X, we could directly compute

$$\mathbb{E}[X^2] = \sum_{i=2}^{6} \frac{i^2(i-1)}{36} + \sum_{i=7}^{12} \frac{i^2(13-i)}{36} = \frac{329}{6}$$

As before, linearity of expectation gives an easier way out. We have,

$$\mathbb{E}[(D_1 + D_2)^2] = \mathbb{E}\left[D_1^2 + D_2^2 + 2D_1D_2\right] \\ = \mathbb{E}[D_1^2] + \mathbb{E}[D_2^2] + 2\mathbb{E}[D_1D_2] \\ = \mathbb{E}[D_1^2] + \mathbb{E}[D_2^2] + 2\mathbb{E}[D_1]\mathbb{E}[D_2],$$

where the last equality follows by the independence of the random variables  $D_1$  and  $D_2$ . We already computed  $\mathbb{E}[D_1] = \mathbb{E}[D_2] = (7/2)$ . Moreover, we also have  $\mathbb{E}[D_1^2] = \mathbb{E}[D_2^2] = \frac{1}{6}\sum_{i=1}^{6} i^2 = \frac{91}{6}$ . Putting everything together, we get

$$\mathbb{E}[X^2] = \frac{91}{6} \times 2 + 2 \times \frac{7}{2} \times \frac{7}{2}$$

$$= \frac{91}{3} + \frac{49}{2} \\ = \frac{329}{6}.$$

Solution Problem 4. (i) Method 1: Direct computation using the probability mass function. Let *B* denote the random variable recording the number of bad apples we have chosen. Let's compute the pmf of *B*. There are  $n^k$  ways to choose *k* apples (with replacement). The number of ways to get  $\ell$  bad apples is  $\binom{k}{\ell} b^{\ell} (n-b)^{k-\ell}$ . Hence, we have

$$\mathbb{P}\left[B=\ell\right] = \frac{\binom{k}{\ell}b^{\ell}(n-b)^{k-\ell}}{n^k}.$$

Therefore,

$$\mathbb{E}[B] = \sum_{\ell=0}^{k} \ell \mathbb{P}[B=\ell]$$
$$= \sum_{\ell=0}^{k} \frac{\ell\binom{k}{\ell} b^{\ell} (n-b)^{k-\ell}}{n^{k}}$$
$$= \sum_{\ell=0}^{k} \ell\binom{k}{\ell} \left(\frac{b}{n}\right)^{\ell} \left(\frac{n-b}{n}\right)^{k-\ell}$$

Here is how to sum this series: From the binomial theorem, we know that

$$(x+y)^k = \sum_{\ell=0}^k \binom{k}{\ell} x^\ell y^{k-\ell}.$$

Differentiating this with respect to x, we see that

$$k(x+y)^{k-1} = \sum_{\ell=1}^{k} \ell \binom{k}{\ell} x^{\ell-1} y^{k-\ell}$$

Hence,

$$kx(x+y)^{k-1} = \sum_{\ell=0}^{k} \ell \binom{k}{\ell} x^{\ell} y^{k-\ell}.$$

Note that this is exactly the expression we got for  $\mathbb{E}[B]$  with  $x = \frac{b}{n}$  and  $y = \frac{n-b}{n}$ . Therefore, we finally get

$$\mathbb{E}[B] = k \frac{b}{n} \left(\frac{b}{n} + \frac{n-b}{n}\right)^{k-1}$$
$$= \frac{kb}{n}.$$

Method 2: Use linearity of expectation.

Let  $X_i$  denote the random variable which is equal to 1 if the  $i^{th}$  apple we choose is bad, and is equal to 0 if the  $i^{th}$  apple we choose is good. Then, the number of bad apples we choose is equal to  $X_1 + \cdots + X_k$  i.e.  $B = X_1 + \cdots + X_k$  Hence, the expected number of bad apples is

$$\mathbb{E}[B] = \mathbb{E}[X_1 + \dots + X_k] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_k].$$

Since we are choosing the apples with replacement,  $\mathbb{E}[X_i] = b/n$  for all *i*. Therefore,

$$\mathbb{E}[B] = \mathbb{E}[X_1 + \dots + X_k] = \frac{kb}{n}.$$

(ii) Method 1: Direct computation using the probability mass function.

Let *B* denote the random variable recording the number of bad apples we have chosen. Let's compute the pmf of *B*. There are  $\binom{n}{k}$  ways to choose *k* apples (without replacement). The number of ways to get  $\ell$  bad apples is  $\binom{b}{\ell}\binom{n-b}{k-\ell}$ . Hence, we have

$$\mathbb{P}\left[B=\ell\right] = \frac{\binom{b}{\ell}\binom{n-b}{k-\ell}}{\binom{n}{k}}.$$

Therefore,

$$\mathbb{E}[B] = \sum_{\ell=0}^{k} \frac{\ell\binom{b}{\ell}\binom{n-b}{k-\ell}}{\binom{n}{k}}.$$

One way to sum this series is the following: first, note that for  $\ell \leq b$ ,  $\ell\binom{b}{\ell} = b\binom{b-1}{\ell-1}$  since

$$\ell \binom{b}{\ell} = \frac{\ell b!}{(b-\ell)!\ell!} = \frac{b(b-1)!}{(b-\ell)!(\ell-1)!} = b\binom{b-1}{\ell-1}.$$

Therefore,

$$\mathbb{E}[B] = \sum_{\ell=0}^{k} \frac{b\binom{b-1}{\ell-1}\binom{n-b}{k-\ell}}{\binom{n}{k}} \\ = b \sum_{\ell=0}^{k} \frac{\binom{b-1}{\ell-1}\binom{n-b}{k-\ell}}{\frac{n}{k}\binom{n-1}{k-1}} \\ = \frac{bk}{n} \sum_{\ell=0}^{k} \frac{\binom{b-1}{\ell-1}\binom{(n-1)-(b-1)}{(k-1)-(\ell-1)}}{\binom{n-1}{k-1}} \\ = \frac{bk}{n}.$$

The last equality holds because  $\sum_{\ell=0}^{k} {\binom{b-1}{\ell-1}} {\binom{(n-1)-(b-1)}{(k-1)-(\ell-1)}} = {\binom{n-1}{k-1}}$  - both sides count the number of ways to choose k-1 apples (without replacement) from a pile of n-1 apples, of which b-1 are bad. Note that this is the same answer as above!

Method 2: Use linearity of expectation.

Again, linearity of expectation is much simpler. Let  $X_i$  denote the random variable which is equal to 1 if the  $i^{th}$  apple we choose is bad, and is equal to 0 if the  $i^{th}$  apple we choose is good. Then, the number of bad apples we choose is equal to  $X_1 + \cdots + X_k$  i.e.  $B = X_1 + \cdots + X_k$  Hence, the expected number of bad apples is

$$\mathbb{E}[B] = \mathbb{E}[X_1 + \dots + X_k] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_k].$$

We will show that it is still the case that  $\mathbb{E}[X_i] = b/n$  for all *i*. In order to see this, note that there are  $\binom{n}{k}k!$  sequences representing the *k* apples chosen (in order). We want to count the number of such sequences which contain a bad apple in position *i*. There are *b* ways to

choose such a bad apple. The remaining k-1 positions can have any of the remaining n-1 apples, hence there are  $\binom{n-1}{k-1}(k-1)!$  such sequences of length k-1. Put together, we see that

$$\mathbb{E}[X_i] = \frac{b\binom{n-1}{k-1}(k-1)!}{\binom{n}{k}k!} \\ = \frac{b(n-1)!/(n-k)!}{n!/\{(n-k)!\}} \\ = \frac{b}{n}.$$

Therefore,

$$\mathbb{E}[B] = \mathbb{E}[X_1 + \dots + X_k] = \frac{kb}{n}.$$

Method 3: Another way of using linearity of expectation.

For  $1 \leq i \leq b$ , let  $Y_i$  be the random variable which is equal to 1 if the  $i^{th}$  bad apple is chosen and 0 otherwise. Then, the number of bad apples chosen is precisely  $B = Y_1 + \cdots + Y_b$ . Therefore, by the linearity of expectation

$$\mathbb{E}[B] = \mathbb{E}[Y_1 + \dots + Y_b] = \mathbb{E}[Y_1] + \dots + \mathbb{E}[Y_b].$$

Finally, note that for all  $1 \leq i \leq b$ ,  $\mathbb{E}[Y_i] = \mathbb{P}[Y_i = 1] = \frac{\binom{n-1}{k-1}}{\binom{n}{k}} = \frac{k}{n}$ , since there are  $\binom{n}{k}$  total ways of choosing k apples, and there are  $\binom{n-1}{k-1}$  ways of choosing k apples such that the  $i^{th}$  bad apple is also chosen. Therefore, as before

$$\mathbb{E}[B] = \mathbb{E}[Y_1 + \dots + Y_b] = \frac{bk}{n}.$$

Solution Optional Problem. Let T denote the number of tosses until the first tail appears. Let W denote your winnings (in dollars). Then,  $W = 2^T$ .

(i) We can easily compute the pmf of T; indeed, for any integer  $i \ge 1$ , we have

$$\mathbb{P}[T=i] = 2^{-i}.$$

Therefore,

$$\mathbb{E}[W] = \sum_{i=1}^{\infty} 2^{i} 2^{-i} = \sum_{i=1}^{\infty} 1 = \infty.$$

(ii) Recall that  $F_W(x) = \mathbb{P}[W \leq x]$ . In our case, note that  $F_W(x) = 0$  for all x < 2. For  $x \geq 2$ , let k denote the largest integer such that  $2^k \leq x$ . Then, we have

$$F_W(x) = \mathbb{P}[W \le x] = \mathbb{P}[T \le k] = \sum_{j=1}^k \mathbb{P}[T=j] = \sum_{j=1}^k 2^{-j} = (1-2^{-k}).$$