STATS 217: Introduction to Stochastic Processes I

Lecture 20

Convergence theorem

- Last time, we proved the convergence theorem for irreducible, aperiodic, finite-state Markov chains.
- Let $(X_n)_{n\geq 0}$ be a DTMC on S with transition matrix P. Suppose that P is irreducible and aperiodic with unique stationary distribution π .
- Let

$$\Delta(n) = \max_{x \in S} \Delta_x(n) = \max_{x \in S} \mathsf{TV}(X_n \mid X_0 = x, \pi).$$

• There exist constants $\epsilon > 0$ and C > 0 (depending on P) such that

$$\Delta(n) \leq C \cdot (1 - \epsilon)^n$$
.

Sub-multiplicativity

• In fact, we worked with the quantities

$$D_{x,y}(n) = TV(X_n \mid X_0 = x, X_n \mid X_0 = y)$$

and

$$D(n) = \max_{x,y \in S} D_{x,y}(n).$$

We showed that

$$\Delta(n) \leq D(n) \leq 2\Delta(n)$$

for all integers $n \ge 0$ and that for any integers $s, t \ge 0$,

$$D(s+t) \leq D(s)D(t).$$

• For $\varepsilon \in [0,1]$, define the ε -mixing time of the chain to be

$$\tau_{\mathsf{mix}}(\varepsilon) := \min\{t : \Delta(t) \le \varepsilon\}.$$

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• The choice of the constant 1/4 is not important and can be replaced by another constant which is strictly smaller than 1/2.

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m mix}$ is because for any $\varepsilon \in (0,1)$,

$$au_{
m mix}(arepsilon) \leq \lceil \log_2 arepsilon^{-1}
ceil au_{
m mix}.$$
 In $arepsilon$, $arepsilon$ = 2

e.g.
$$e = 2^{-100}$$
 $T_{mix}(2^{-100}) \leq 100 T_{mix}(Y_4)$

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$$\tau_{\mathsf{mix}}(\varepsilon) \leq \lceil \log_2 \varepsilon^{-1} \rceil \tau_{\mathsf{mix}}.$$

Indeed,

$$\sum_{(n)} \leq \sum_{(n)} \leq 2\Delta \\
(n) \leq (n)$$

$$\Delta(\lceil \log_2 \varepsilon^{-1} \rceil \tau_{\text{mix}}) \leq D(\lceil \log_2 \varepsilon^{-1} \rceil \cdot \tau_{\text{mix}})$$

$$\leq D(\tau_{\text{mix}})^{\lceil \log_2 \varepsilon^{-1} \rceil}$$

$$\int (S \downarrow \leftarrow) \leq \int (S)^{\beta} (c)$$

$$\int (K \downarrow S) \leq \int (S)^{\kappa}$$

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Indeed,

$$\begin{split} \Delta(\lceil \log_2 \varepsilon^{-1} \rceil \tau_{\mathsf{mix}}) &\leq D(\lceil \log_2 \varepsilon^{-1} \rceil \cdot \tau_{\mathsf{mix}}) \\ &\leq D(\tau_{\mathsf{mix}})^{\lceil \log_2 \varepsilon^{-1} \rceil} \\ &\leq (2\Delta(\tau_{\mathsf{mix}}))^{\lceil \log_2 \varepsilon^{-1} \rceil} \\ &\triangleq (\mathsf{Z}_{\mathsf{M}}(\mathsf{x})) &\leq \mathsf{V}_{\mathsf{H}} \\ &=) \quad 2 \, \mathsf{A} \, (\mathsf{T}_{\mathsf{M}}(\mathsf{x})) &\leq \mathsf{V}_{\mathsf{L}} \end{split}$$

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 $\Delta(\lceil \log_2 \varepsilon^{-1} \rceil \tau_{\mathsf{mix}}) \leq D(\lceil \log_2 \varepsilon^{-1} \rceil \cdot \tau_{\mathsf{mix}}) \qquad \text{ for smaller.}$

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Indeed,

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- Consider a transition matrix P on a finite state space S.
- A coupling of Markov chains with transition matrix P and initial distributions μ and ν is a process

$$(\widehat{X}_t, \widehat{Y}_t)_{t=0}^{\infty}$$

such that for all $t \geq 0$,

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$$\widehat{X}_t \sim (X_t \mid X_0 \sim \mu)$$

 $\widehat{Y}_t \sim (X_t \mid X_0 \sim \nu)$

and such that

$$\widehat{X}_t = \widehat{Y}_t \implies \widehat{X}_{t+1} = \widehat{Y}_{t+1}.$$
 So so the form $\widehat{X}_t = \widehat{Y}_t = \widehat{Y}_t = \widehat{Y}_t$ reasonable "couplings.

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and such that

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• We have already seen couplings of Markov chains in our proof of the convergence theorem in this case, wh used

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to guarantee the existence of couplings of certain properties.

• As we will soon see, couplings of Markov chains are a useful tool to bound the mixing time in applications.

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- This is due to the following: Let $(\widehat{X}_t, \widehat{Y}_t)$ be a coupling of two Markov chains with transition matrix P and with $\widehat{X}_0 = x$, $\widehat{Y}_0 = y$. Let

$$\tau_{\text{couple}} := \min\{t : \widehat{X}_t = \widehat{Y}_t\}.$$

$$\stackrel{\wedge}{\chi}_t = \stackrel{\wedge}{Y}_t \implies \stackrel{\wedge}{\chi}_{t+1} = \stackrel{\wedge}{Y}_{t+1}$$

$$\text{we know that}$$

$$t > \tau_{\text{couple}} := \min\{t : \widehat{X}_t = \widehat{Y}_t\}.$$

- As we will soon see, couplings of Markov chains are a useful tool to bound the mixing time in applications.
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$$au_{\mathsf{couple}} := \min\{t : \widehat{X}_t = \widehat{Y}_t\}.$$

Recall that

$$D_{x,y}(n) = \text{TV}(X_n \mid X_0 = x, X_n \mid X_0 = y)$$
. Coupling for which the dis-

Then,

$$D_{x,y}(n) \leq \mathbb{P}[au_{\mathsf{couple}} \geq n].$$
 of $\mathsf{T}_{\mathsf{couple}}$ can be studied.

construct a

not - too - hard

$$R$$
 then use $\triangle(n) \leq D(n) = \max_{x_i, y_i} D(x_i, y_i)$.

- The proof is a direct application of the coupling lemma.
- Indeed, since $\widehat{X}_n \sim X_n \mid X_0 = x$ and $\widehat{Y}_n \sim X_n \mid X_0 = y$, we have

$$D_{x,y}(n) \leq \mathbb{P}[\widehat{X}_n \neq \widehat{Y}_n] \leq \mathbb{P}[\tau_{\text{couple}} \geq n].$$

$$\tau \vee (x_n \mid x_0 = x, x_n \mid x_0 = y)$$

$$\tau \vee (\widehat{x}_n, \widehat{Y}_n) \leq \mathbb{P}[\widehat{x}_n \neq \widehat{Y}_n]$$

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$$D_{x,y}(n) \leq \mathbb{P}[\widehat{X}_n \neq \widehat{Y}_n] \leq \mathbb{P}[\tau_{\text{couple}} \geq n].$$

Therefore, by Markov's inequality,

$$D_{x,y}(4 \cdot \mathbb{E}[au_{\text{couple}}]) \leq \mathbb{P}[au_{\text{couple}} \geq 4 \cdot \mathbb{E}[au_{\text{couple}}]] \leq rac{1}{4}.$$
 They im $M = H \cdot 1 E \subset C_{\text{couple}}$

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- $S = \{0, 1\}^n$.
- The transitions are given as follows. Suppose the current state is x. With probability 1/2, the chain remains at x; with probability 1/2, it moves uniformly to one of the n possible vectors y which differ from x in exactly one coordinate.
- The transition matrix is clearly aperiodic and irreducible, and we have seen that the unique stationary distribution is the uniform distribution on $\{0,1\}^n$.

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- The transition matrix is clearly aperiodic and irreducible, and we have seen that the unique stationary distribution is the uniform distribution on $\{0,1\}^n$.
- Here is an equivalent description of the transitions: suppose the current state is x. We choose a coordinate $i \in \{1, \ldots, n\}$ uniformly at random and an unbiased bit $b \in \{0, 1\}$, also uniformly at random, and independent of the coordinate i.

• Given this alternate description, there is a natural choice of coupling: for the two chains started from x and y, use the same i and b at every step.

$$\hat{x}_{0} = x \qquad \hat{y}_{0} = y$$

$$\hat{y}_{1} = y$$

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$$\hat{y}_{1} = y$$

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$$\hat{y}_{4} = y$$

$$\hat{y}$$

- Given this alternate description, there is a natural choice of coupling: for the two chains started from x and y, use the same i and b at every step.
- Let au denote the first time that each coordinate i has been chosen to be updated. Then, clearly, $\widehat{X}_{\tau} = \widehat{Y}_{\tau}$.

when we update coord i
$$(\hat{\chi}_{\xi+1})_{i} = b = (\hat{\chi}_{\xi+1})$$

- Given this alternate description, there is a natural choice of coupling: for the two chains started from x and y, use the same i and b at every step.
- Let au denote the first time that each coordinate i has been chosen to be updated. Then, clearly, $\widehat{X}_{\tau} = \widehat{Y}_{\tau}$.
- Moreover, τ is exactly the first time to collect all n coupons in the coupon-collector problem and

poblem and we have a coupons
$$\mathbb{P}[\tau>t] \leq \overline{n} \bigg(1-\frac{1}{n}\bigg)^t \leq ne^{-t/n},$$

$$\text{projethat}$$
 union coupon in has not lead by home to

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$$\mathbb{P}[\tau > t] \le n \left(1 - \frac{1}{n}\right)^t \le n e^{-t/n},$$

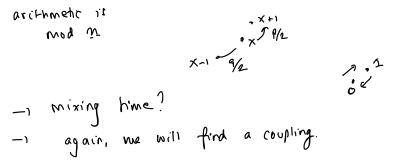
which gives $\tau_{\text{mix}} \leq n \log n + n \log(1/4)$.

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• The states of the *n*-cycle can be identified with \mathbb{Z}_n , the integers modulo *n*.

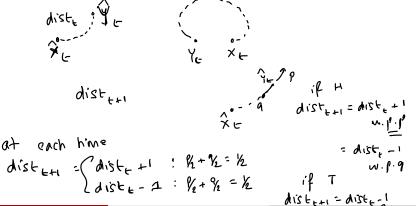
- The states of the *n*-cycle can be identified with \mathbb{Z}_n , the integers modulo *n*.
- The transitions are given as follows. Suppose that the current state is x. With probability 1/2, the chain remains at x; with probability p/2, it moves to x+1; with probability q/2, it moves to x-1. Here, p+q=1.



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- The transitions are given as follows. Suppose that the current state is x. With probability 1/2, the chain remains at x; with probability p/2, it moves to x + 1; with probability q/2, it moves to x - 1. Here, p + q = 1.
- Here is a natural choice of coupling: start the two chains at x and y. At each step, flip an unbiased coin. If the coin lands heads, then the x-chain stays put, and the y chain moves +1 with probability p and -1 with probability q. If the coin lands tails, then the y-chain stays put, and the x chain moves +1with probability p and -1 with probability q.
- consider staying nhere you are: 1/2
 the x-chain = 1 : 9/2

• Let $dist_t$ denote the (clockwise) distance between the states of the two chains at time t.

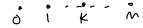


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- Let dist_t denote the (clockwise) distance between the states of the two chains at time t.
- Then, $(\operatorname{dist}_t)_{t\geq 0}$ is a simple symmetric random walk on $\{0,\ldots,n\}$ with absorbing states 0 and n.



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- Then, $(\operatorname{dist}_t)_{t\geq 0}$ is a simple symmetric random walk on $\{0,\ldots,n\}$ with absorbing states 0 and n.
- From Gambler's ruin, we know that if the initial distance is k, then the expected time to absorption is $k(n-k) \le n^2/4$.
- Hence,

$$\tau_{\rm mix} \leq 4 \cdot \frac{n^2}{4} = n^2.$$
 is this a good bound?
$$y_{\rm RS_r} = v_{\rm p} + t_{\rm o} = constant$$