STATS 217: Introduction to Stochastic Processes I

Lecture 21

Lectu

(Diaconis)

Stationary times

• Last time, we used the following description of the lazy random walk on the hypercube: suppose the current state is x and the current time is t. We choose a coordinate $i_{t+1} \in \{1, \ldots, n\}$ uniformly at random and an unbiased bit $b_{t+1} \in \{0, 1\}$, also uniformly at random, and independently of the coordinate i_{t+1} . Then, we set the value of coordinate i_{t+1} to b_{t+1} and keep all other coordinates unchanged.

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- Let $\tau_{\rm refresh}$ be the first time that all coordinates have been chosen to be updated.
- Note that τ_{refresh} is a stopping time with respect to the collection of random tuples {(*i_k*, *b_k*)}_{k≥1}. (Why?)

stopping time with Xt

Stationary times m= 8 000

$$i_1 = 1$$
 $b_1 = 0$
 $i_2 = 2$ $b_2 = 0$
 $i_3 = 3$ $b_3 = 0$

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- Let $\tau_{\rm refresh}$ be the first time that all coordinates have been chosen to be updated.
- Note that τ_{refresh} is a stopping time with respect to the collection of random tuples $\{(i_k, b_k)\}_{k \ge 1}$. (Why?)

• Moreover, note that $X_{\tau_{\text{refresh}}}$ is distributed uniformly on $\{0,1\}^n$. (Why?)



- τ_{refresh} is an example of a stationary time.
- By this, we mean a stopping time τ (with respect to the initial state x of the chain as well as any auxiliary randomness) which satisfies

$$\mathbb{P}[X_{ au} = y \mid X_0 = x] = \pi(y) \quad \forall y \in S,$$

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• Let $x \in S$ and consider the chain with initial state $X_0 = x$. A stationary time τ for which τ and X_{τ} are independent i.e. for all integers t and $y \in S$

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• τ_{refresh} is an example of a strong stationary time for any starting state x. (Why?)

• A strong stationary time for the starting state x is a stopping time (with respect to the chain and auxiliary randomness) satisfying

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$$\bigcup + \text{Iaw of total probability}$$

• The property that will be useful to us is

$$\left[\mathbb{P}[\tau \leq t, X_t = y \mid X_0 = x] = \mathbb{P}[\tau \leq t \mid X_0 = x] \cdot \pi(y) \right]$$

which can be proved using the law of total probability.

$$\mathbb{P}[\tau \le t, X_t = y \mid X_0 = x] = \sum_{s \le t} \sum_{z \in S} \mathbb{P}[\tau = s, X_t = y, X_s = z \mid X_0 = x]$$

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$$TV(X_{t} | X_{0} = x, \pi) \leq \mathbb{P}[\tau > t | X_{0} = x].$$

$$\frac{1}{2} \sum_{y=\pi}^{1} [\pi(y) - p_{x,y}^{t}] \quad \forall \quad \forall \quad T \cup (Ber(p), Ber(p))$$

$$= \frac{1}{2} \ln(p) + \frac{1}{2} \ln$$

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- Strong stationary times are also a useful tool for bounding the mixing time. This is captured by the following:
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$$\sum_{x \in Y} \pi(y) \leq 1$$

- Strong stationary times are also a useful tool for bounding the mixing time. This is captured by the following:
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Note that

$$TV(X_t \mid X_0 = x, \pi) = \sum_{\substack{y:\pi(y) > P_{x,y}^t \\ y:\pi(y) > P_{x,y}^t}} [\pi(y) - P_{x,y}^t]$$
$$= \sum_{\substack{y:\pi(y) > P_{x,y}^t \\ y \in S}} \pi(y) \left(1 - \frac{P_{x,y}^t}{\pi(y)}\right)$$
$$\leq \max_{\substack{y \in S}} \left(1 - \frac{P_{x,y}^t}{\pi(y)}\right).$$

• Hence, it will suffice to show that for all $y \in S$,

$$1 - \frac{P_{x,y}^{t}}{\pi(y)} \leq \mathbb{P}[\tau > t \mid X_{0} = x].$$

$$\Rightarrow \quad \forall v (p^{t}(x_{i}, \cdot), \pi) \leq \max_{\substack{y \in S}} \left(1 - \frac{p_{x_{i}}}{\pi(y)} \right)$$

$$\leq \mathbb{P}\left[\forall \mathcal{F}_{0} \in [x_{0} = x] \right].$$

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• We have

$$1 - \frac{P_{x,y}^{t}}{\pi(y)} = 1 - \frac{\mathbb{P}[X_{t} = y \mid X_{0} = x]}{\pi(y)} \int \\ \leq 1 - \frac{\mathbb{P}[X_{t} = y, \tau \leq t \mid X_{0} = x]}{\pi(y)} \\ = 1 - \frac{\pi(y) \mathbb{P}[\tau \leq t \mid X_{0} = x]}{\pi(y)} \\ = \frac{\pi(y) \mathbb{P}[\tau \leq t \mid X_{0} = x]}{\pi(y)}$$

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$$\begin{aligned} 1 - \frac{P_{x,y}^{t}}{\pi(y)} &= 1 - \frac{\mathbb{P}[X_{t} = y \mid X_{0} = x]}{\pi(y)} \\ &\leq 1 - \frac{\mathbb{P}[X_{t} = y, \tau \leq t \mid X_{0} = x]}{\pi(y)} \\ &= 1 - \frac{\pi(y)\mathbb{P}[\tau \leq t \mid X_{0} = x]}{\pi(y)} \end{aligned}$$

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$$\begin{cases} x & \text{shong hmes} \\ \forall \forall \forall (p^{t}(x_{j}, t), \pi) \\ \leq \mathbb{P}[\tau > t \mid X_{0} = x]. \end{cases}$$

• We have

$$\begin{split} 1 - \frac{P_{x,y}^t}{\pi(y)} &= 1 - \frac{\mathbb{P}[X_t = y \mid X_0 = x]}{\pi(y)} & \text{find a not} \\ &\leq 1 - \frac{\mathbb{P}[X_t = y, \tau \leq t \mid X_0 = x]}{\pi(y)} & z \\ &= 1 - \frac{\pi(y)\mathbb{P}[\tau \leq t \mid X_0 = x]}{\pi(y)} & z \\ &= \mathbb{P}[\tau > t \mid X_0 = x]. \end{split}$$

$$\begin{array}{c} A \longrightarrow \\ B \longrightarrow \\ C \longrightarrow \\$$

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of permutation of
$$\{1, ..., n\}$$

= $n!$
 $\sim ... \left(\frac{n}{e}\right)^n$

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- At each step of the top-to-random shuffle, we take the top card, and place it in any of the *n* available positions.
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- Now, we will bound the mixing time by constructing a suitable strong stationary time.

 \bullet Let $\tau_{\rm top}$ be first time when the original bottom card has moved to the top.

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- Let τ_{top} be first time when the original bottom card has moved to the top. • Then, $\tau = \tau_{top} + 1$ is a strong stationary time.
 - Xo-A B.----X_{Cop} B now G A select A D C D A C

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- The key to showing that it is a strong stationary time is the following observation, which can be proved by induction: let τ_k denote the first time when there are k cards under the original bottom card. Then, at X_{τ_k} , all k! permutations of these cards are equally likely. (Why?)



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- The key to showing that it is a strong stationary time is the following observation, which can be proved by induction: let τ_k denote the first time when there are k cards under the original bottom card. Then, at X_{τ_k} , all k! permutations of these cards are equally likely. (Why?)
- Therefore, by the relationship between strong stationary times and total variation distance,

$$\mathsf{TV}(X_t \mid X_0 = x, \pi) \leq \mathbb{P}[\tau > t \mid X_0 = x].$$

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- T = Ttop + 1 first hime when all cards are lelow ofiginal bottom card.
- Note that, when there are k cards under the bottom card, the probability that the top card goes under it is (k + 1)/n.
- Therefore, we see that

$$\tau \sim \operatorname{Geom}(1/n) + \operatorname{Geom}(2/n) + \cdots + \operatorname{Geom}(1),$$

where the geometric random variables on the right hand side are independent. This is the same as the distribution of the coupon collector's time.

• Therefore, as for the lazy random walk on the hypercube, we get that