

this week: martingales

STATS 217: Introduction to Stochastic Processes I

Lecture 26

Martingales

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- We say that the sequence of random variables M_0, M_1, \dots is a **martingale** with respect to X_1, X_2, \dots if

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- $\mathbb{E}[M_n | X_1, X_2, \dots, X_{n-1}] = M_{n-1}$. Explicitly, for any x_1, \dots, x_{n-1} ,

$$\mathbb{E}[M_n | X_1 = x_1, \dots, X_{n-1} = x_{n-1}] = f_n(x_1, \dots, x_{n-1}).$$

$$\begin{aligned} (\Leftrightarrow) \{ & \mathbb{E}[M_n - M_{n-1} | X_1, \dots, X_{n-1}] = 0 \} \\ & \Leftrightarrow \mathbb{E}[M_n | X_1, \dots, X_{n-1}] = \mathbb{E}[M_{n-1} | X_1, \dots, X_{n-1}] \end{aligned}$$

Example

$$\begin{aligned} \mathbb{E}(M_{n-1} | X_1 = x_1, \dots, X_{n-1} = x_{n-1}) &= f_{n-1}(x_1, \dots, x_{n-1}) \\ &= M_{n-1} \end{aligned}$$

e.g. simple symmetric r.w. in 1-D

$$\text{then } X_i = \begin{cases} +1 & \text{w.p. } 1/2 \\ -1 & \text{w.p. } 1/2 \end{cases}$$

- X_1, X_2, \dots are independent random variables with $\mathbb{E}[X_i] = 0$ for all $i \geq 1$.
- Let $M_0 = 0$ and for $n \geq 1$,

$$\mathbb{E}[|X_i|] < \infty$$

$$M_n = X_1 + \dots + X_n.$$

Example

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- Then, M_0, M_1, \dots is a martingale with respect to X_1, X_2, \dots .

$$(1) \quad \mathbb{E}[|M_n|] \leq \mathbb{E}[|X_1|] + \dots + \mathbb{E}[|X_n|] < \infty.$$

$$(2) \quad M_n = f(X_1, \dots, X_n)$$

$$(3) \quad \mathbb{E}[M_n | X_1, \dots, X_{n-1}] = M_{n-1}$$

$$\begin{aligned} (\Rightarrow) \quad \mathbb{E}[M_n - M_{n-1} | X_1, \dots, X_{n-1}] &= \mathbb{E}[X_n | X_1, \dots, X_{n-1}] \\ &= \mathbb{E}[X_n] = 0. \end{aligned}$$

Example

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- Let $M_0 = 0$ and for $n \geq 1$,

$$M_n = X_1 + \dots + X_n.$$

- Then, M_0, M_1, \dots is a martingale with respect to X_1, X_2, \dots .
- This generalizes the one-dimensional simple, symmetric random walk.

Example

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- Then, $M_0 = 0$ and for $n \geq 1$,

$$M_n = (X_1 + \dots + X_n)^2 - n\sigma^2 = S_n^2 - n\sigma^2$$

is a martingale with respect to X_1, X_2, \dots

$$* \quad \underline{M}_n = f_n(X_1, \dots, X_n)$$

$$* \quad \mathbb{E}[\underline{M}_n - \underline{M}_{n-1} \mid X_1, \dots, X_{n-1}] = 0.$$

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$$M_n = (X_1 + \dots + X_n)^2 - n\sigma^2$$

is a martingale with respect to X_1, X_2, \dots .

- To verify the martingale property, note that

$$\begin{aligned} \mathbb{E}[M_n - M_{n-1} \mid X_1, \dots, X_{n-1}] &= \mathbb{E}[\underbrace{(X_n + S_{n-1})^2}_{\text{square}} - \underbrace{S_{n-1}^2}_{\text{square}} - \underbrace{\sigma^2}_{\text{constant}} \mid X_1, \dots, X_{n-1}] \\ &= \mathbb{E}[(X_n + S_{n-1})^2 - S_{n-1}^2 - \sigma^2 \mid X_1, \dots, X_{n-1}] \\ &= \mathbb{E}[X_n^2 + 2X_n S_{n-1} + S_{n-1}^2 - S_{n-1}^2 - \sigma^2 \mid X_1, \dots, X_{n-1}] \\ &= \mathbb{E}[X_n^2 + 2X_n S_{n-1} - \sigma^2 \mid X_1, \dots, X_{n-1}] \\ &= \mathbb{E}[X_n^2 - \sigma^2 + 2X_n S_{n-1} \mid X_1, \dots, X_{n-1}] \\ &= \mathbb{E}[X_n^2 - \sigma^2] + 2S_{n-1} \mathbb{E}[X_n \mid X_1, \dots, X_{n-1}] \\ &= \sigma^2 - \sigma^2 + 2S_{n-1} \cdot 0 \\ &= 0 \end{aligned}$$

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$$\begin{aligned}&= \underbrace{\mathbb{E}[X_n^2 \mid X_1, \dots, X_{n-1}]}_{\substack{\text{"} \\ \mathbb{E}[X_n^2] = \text{Var}(X_n) \\ = \sigma^2}} - \cancel{\sigma^2} \\ &\quad + 2 \mathbb{E}[X_n S_{n-1} \mid X_1, \dots, X_{n-1}] \\ &= 2 \underbrace{S_{n-1}}_{\substack{\text{"} \\ \mathbb{E}[X_n \mid X_1, \dots, X_{n-1}]}}\end{aligned}$$

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$$= 2S_{n-1} \mathbb{E}[X_n] = 0.$$

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- Then, $M_0 = 1$ and for $n \geq 1$,

$$M_n = M_0 \cdot X_1 \cdots X_n$$

is a martingale with respect to X_1, \dots, X_n .

$$\begin{aligned} & \mathbb{E} [M_n \mid X_1, \dots, X_{n-1}] \\ &= \mathbb{E} [M_{n-1} X_n \mid X_1, \dots, X_{n-1}] \\ &= M_{n-1} \mathbb{E} [X_n] = 1 \cdot M_{n-1} = M_{n-1} \end{aligned}$$

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- Let $X_i = e^{\lambda Y_i} / \phi(\lambda)$. Then, X_1, X_2, \dots are independent random variables with $\mathbb{E}[X_i] = 1$.

$$\frac{\mathbb{E}[e^{\lambda Y_i}]}{\phi(\lambda)} = \frac{\phi(\lambda)}{\phi(\lambda)} = 1.$$

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- Let $X_i = e^{\lambda Y_i} / \phi(\lambda)$. Then, X_1, X_2, \dots are independent random variables with $\mathbb{E}[X_i] = 1$.
- Therefore, $M_0 = 1$ and for $n \geq 1$,

$$M_n = M_0 \cdot X_1 \cdots X_n = e^{\lambda(Y_1 + \cdots + Y_n)} / \phi(\lambda)^n$$

is a martingale with respect to Y_1, Y_2, \dots .

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- Recall this means that

$$Z_n = \sum_{i=1}^{Z_{n-1}} \xi_i, \quad \mathbb{E}[M_n | Z_0, \dots, Z_{n-1}]$$
$$= \mathbb{E}[M_{n-1} | Z_{n-1}]$$

where ξ_1, ξ_2, \dots are i.i.d. copies of ξ .

- The sequence $M_0 = 1$ and for $n \geq 1$,

$$M_n = \frac{Z_n}{\mu^n}$$

is a martingale with respect to M_1, M_2, \dots

$$= \frac{1}{\mu^n} \mathbb{E}[Z_n | Z_{n-1}]$$
$$= \frac{1}{\mu^n} [\mu \cdot Z_{n-1}]$$
$$= \frac{Z_{n-1}}{\mu^{n-1}} = M_{n-1}$$

Submartingales and supermartingales

- A **supermartingale** is defined similarly to a martingale, except now we weaken the martingale condition to

$$\mathbb{E}[M_n \mid X_1, \dots, X_{n-1}] \leq M_{n-1}.$$

\Leftrightarrow for m.g.

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- On the other hand, **submartingales** are favorable to the gambler i.e., they satisfy

$$\mathbb{E}[M_n \mid X_1, \dots, X_{n-1}] \geq M_{n-1}.$$

Martingale betting strategy

in the background: ind. fair
coin
tosses

- Consider a gambling game based on successive outcomes of a fair coin toss.
- You adopt the following strategy: if you win a round, then in the next round, you bet \$1; if you lose a round, then in the next round, you double your bet.

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- So, for instance, if you lose in the first three rounds, and win in the fourth round, your sequence of bets is \$1, \$2, \$4, \$8, and your net winnings are

$$-\$1 - \$2 - \$4 + \$8 = 1.$$

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- More generally, if you lose the first  $k$  rounds and win the  $k + 1^{\text{st}}$  round, your net winnings are

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- Moreover, in an infinite sequence of fair coin tosses, you will win with probability 1.

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- Then, your net winnings are:

|       |                               |   |      |
|-------|-------------------------------|---|------|
| $TTT$ | $-\$7$                        | } | $-6$ |
| $TTH$ | $+\$1$                        |   |      |
| $THT$ | $+\$0$                        |   |      |
| $THH$ | <del><math>+\\$2</math></del> |   |      |
| $HTT$ | <del><math>-\\$2</math></del> |   |      |
| $HTH$ | $+\$2$                        | } | $+6$ |
| $HHT$ | $+\$1$                        |   |      |
| $HHH$ | $+\$3$                        |   |      |
|       | <hr/>                         |   |      |
|       | $0$                           |   |      |



## Martingale betting strategy

- Let's take a look at this game for a fixed number of rounds, say 3 rounds. Suppose a win for you corresponds to  $H$ .
- Then, your net winnings are:

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$TTH$      $+ \$1$

$THT$      $+ \$0$

$THH$      $+ \$2$

$HTT$      $- \$2$

$HTH$      $+ \$2$

$HHT$      $+ \$1$

$HHH$      $+ \$3$

- Therefore, if  $M_3$  denotes your winnings after 3 rounds of the game using the martingale betting strategy, then

$$\mathbb{E}[M_3] = 0.$$

# Martingale transforms

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- A sequence of random variables  $A_1, A_2, \dots$  is called predictable with respect to the sequence  $X_1, X_2, \dots$  if for all  $n \geq 1$ ,

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- So, if we think of  $X_1, X_2, \dots$  as being the outcomes of rounds of a gambling game, then  $A_n$  is a function of the information that the gambler has *before* placing the bet in the  $n^{\text{th}}$  round.

# Martingale transforms

- Let  $M_0, M_1, \dots$  be a martingale with respect to  $X_1, X_2, \dots$ , and let  $A_1, A_2, \dots$  be a predictable sequence with respect to  $X_1, X_2, \dots$ .



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- The **martingale transform** of  $\{M_n\}$  by  $\{A_n\}$  is defined by  $\tilde{M}_0 = M_0$  and for  $n \geq 1$ ,

$$\tilde{M}_n = M_0 + A_1(M_1 - M_0) + A_2(M_2 - M_1) + \dots + A_n(M_n - M_{n-1}).$$

  $\rightarrow$  martingale differences

- Intuition:  $(M_k - M_{k-1})$  is the gain from the  $k^{\text{th}}$  round of the gambling game. The gambler looks at all previous outcomes  $X_1, \dots, X_{k-1}$ , and comes up with a multiplier  $A_k$  for the  $k^{\text{th}}$  round.



## Martingale transforms are martingales

- Let  $M_0, M_1, \dots$  be a martingale with respect to  $X_1, X_2, \dots$ , and let  $A_1, A_2, \dots$  be a predictable sequence with respect to  $X_1, X_2, \dots$ .
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- Let  $\tilde{M}_0, \tilde{M}_1, \dots$  be the martingale transform of  $\{M_n\}$  by  $\{A_n\}$ .
- Then,  $\tilde{M}_0, \tilde{M}_1, \dots$  is also a martingale with respect to  $X_1, X_2, \dots$ .

$$\tilde{M}_n = f_n(X_1, \dots, X_n)$$

$$\mathbb{E}[\tilde{M}_n - \tilde{M}_{n-1} \mid X_1, \dots, X_{n-1}] = 0.$$

$$\tilde{M}_n - \tilde{M}_{n-1} = A_n \cdot (M_n - M_{n-1})$$

$$\begin{aligned}\mathbb{E}[A_n \cdot (M_n - M_{n-1}) \mid X_1, \dots, X_{n-1}] &= A_n \mathbb{E}[M_n - M_{n-1} \mid X_1, \dots, X_{n-1}] \\ &= A_n \cdot 0 = 0\end{aligned}$$

## Martingale transforms are martingales

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- Let  $\tilde{M}_0, \tilde{M}_1, \dots$  be the martingale transform of  $\{M_n\}$  by  $\{A_n\}$ .
- Then,  $\tilde{M}_0, \tilde{M}_1, \dots$  is also a martingale with respect to  $X_1, X_2, \dots$ .
- Indeed,

$$\begin{aligned}\mathbb{E}[\tilde{M}_n - \tilde{M}_{n-1} \mid X_1, \dots, X_{n-1}] &= \mathbb{E}[A_n(M_n - M_{n-1}) \mid X_1, \dots, X_{n-1}] \\ &= A_n \cdot \mathbb{E}[M_n - M_{n-1} \mid X_1, \dots, X_{n-1}] \\ &= 0.\end{aligned}$$